# Fast optimal bandwidth selection for kernel density estimation

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## Gist of the paper

- Bandwidth selection for kernel density estimation scales as  $\mathcal{O}(N^2)$ .
- We present a fast computational technique that scales as  $\mathcal{O}(N)$ .
- For 50,000 points we obtained speedups in the range 65 to 105.

#### **Density estimation**

- Widely used in exploratory data analysis, machine learning, data mining, computer vision, and pattern recognition.
- A density p gives a principled way to compute probabilities on sets.

$$\Pr[x \in A] = \int_A p(x) dx.$$

• Estimate the density from samples  $x_1, \ldots, x_N$  drawn from p.

# Different methods for density estimation

- Parametric methods.
  - Assume a functional form for the density.
- Non-parametric methods.
  - letting the data speak for themselves
  - Histograms.
  - Kernel density estimators. [Most popular.]

Kernel density estimate (KDE)

$$\widehat{p}(x) = \frac{1}{Nh} \sum_{i=1}^{N} K\left(\frac{x - x_i}{h}\right)$$

- The kernel function K is essentially spreading a probability mass of 1/N associated with each point about its neighborhood.
- The neighborhood size is essentially decided by the parameter *h* called the **bandwidth** of the kernel.

## **KDE** illustration



#### Gaussian kernel

The most widely used kernel is the Gaussian of zero mean and unit variance.

$$K(u) = \frac{1}{\sqrt{2\pi}}e^{-u^2/2}.$$

For the Gaussian kernel the kernel density estimate can be written as

$$\widehat{p}(x) = \frac{1}{N\sqrt{2\pi h^2}} \sum_{i=1}^{N} e^{-(x-x_i)^2/2h^2}.$$

# Computational complexity of KDE

- Essentially sum of N Gaussians.
- Computing KDE at M points scales as  $\mathcal{O}(MN)$ .
- Various approximate algorithms are proposed that reduce the computational complexity to  $\mathcal{O}(M+N)$ .

[FFT, FGT, IFGT, dual-tree].

• This paper focuses on reducing the computational complexity of finding the optimal bandwidth, which scales as  $\mathcal{O}(N^2)$ .

#### Role of bandwidth h

- As h decreases towards 0, the number of modes increases to the number of data points and the KDE is very noisy.
- As h increases towards ∞, the number of modes drops to 1, so that any interesting structure has been smeared away and the KDE just displays a unimodal pattern.



The bandwidth h has to be chosen optimally.



The sense in which the bandwidth is optimal has to be made precise.

The most widely used is the AMISE optimal bandwidth.

#### **Performance measure**

• Integrated squared error (ISE)

$$\mathsf{ISE}(\widehat{p}, p) = \int_{\mathbf{R}} [\widehat{p}(x) - p(x)]^2 dx.$$

• Mean integrated squared error (MISE)

$$\mathsf{MISE}(\widehat{p}, p) = E[\mathsf{ISE}(\widehat{p}, p)] = E\left[\int_{\mathbf{R}} [\widehat{p}(x) - p(x)]^2 dx\right]$$

• A measure of the 'average' performance of the kernel density estimator, averaged over the support of the density and different realization of the points.

## Asymptotic performance measure

- The dependence of the MISE on the bandwidth *h* is not very explicit.
- This makes it difficult to interpret the influence of the bandwidth on the performance of the estimator.
- An asymptotic large sample approximation for this expression is usually derived via the Taylor's series called as the AMISE, the A is for asymptotic.

#### AMISE

The AMISE can be shown to be  $^{\ast}$ 

AMISE
$$(\hat{p}, p) = \frac{1}{Nh}R(K) + \frac{1}{4}h^4\mu_2(K)^2R(p'')$$
  
= Variance + (bias)<sup>2</sup>

where

$$R(g) = \int_{\mathbf{R}} g(x)^2 dx$$
, ,  $\mu_2(g) = \int_{\mathbf{R}} x^2 g(x) dx$ ,

\*Wand, M. P. and Jones, M. C. 1995. Kernel Smoothing. Chapman and Hall, London.

## **Bias-Variance tradeoff**



- Variance is proportional to 1/h.
- Bias is proportional to  $h^2$ .
- Optimal h is found by setting the first derivative of AMISE to zero.

## AMISE optimal bandwidth

$$h_{optimal} = \left[\frac{R(K)}{\mu_2(K)^2 R(p'')N}\right]^{1/5}.$$

- This expression cannot be used directly since R(p'') depends on the second derivative of the density p.
- Different strategies have been proposed to solve this problem.
- The most popular **plug-in methods** use an estimate of R(p'') which in turn needs an estimate of p''.
- So for optimal bandwidth estimation we need estimates of the density derivatives.

# Estimating density functionals

- This bandwidth for estimation of the density functional R(p'') is quite different from the the bandwidth h used for the kernel density estimate.
- We can find an expression for the optimal bandwidth for the estimation of R(p'').
- However this bandwidth will depend on an unknown density functional  $R(p^{'''})$ .
- This problem will continue since the optimal bandwidth for estimating  $R(p^{(s)})$  will depend on  $R(p^{(s+1)})$ .
- In general we need estimates of higher order derivatives also.

#### Kernel density derivative estimation

• Take the derivative of the kernel density estimate.

$$\hat{p}^{(r)}(x) = \frac{1}{Nh^{r+1}} \sum_{i=1}^{N} K^{(r)}\left(\frac{x-x_i}{h}\right).$$

• For the Gaussian kernel this takes the form

$$\hat{p}^{(r)}(x) = \frac{(-1)^r}{\sqrt{2\pi}Nh^{r+1}} \sum_{i=1}^N H_r\left(\frac{x-x_i}{h}\right) e^{-(x-x_i)^2/2h^2}.$$

•  $H_r(u)$  is the  $r^{th}$  Hermite polynomial.

# Computational complexity of bandwidth estimation

- In order to estimate a density functional we need to evaluate the density derivative at N points.
- Hence computing a density functional is  $\mathcal{O}(rN^2)$ .
- The current most practically successful approach, solve-theequation plug-in \* method involves the numerical solution of a non-linear equation.
  - Iterative methods to solve this equation will involve repeated use of the density derivative functional estimator for different bandwidths which adds much further to the computational burden.

\*Sheather, S. and Jones, M. 1991. A reliable data-based bandwidth selection method for kernel density estimation. Journal of Royal Statistical Society Series B 53, 683-690.

Fast  $\epsilon - exact$  density derivative estimation

$$G_r(y_j) = \sum_{i=1}^N q_i H_r\left(\frac{y_j - x_i}{h}\right) e^{-(y_j - x_i)^2/2h^2} \quad j = 1, \dots, M,$$

- The computational complexity is  $\mathcal{O}(rNM)$ .
- We will present an  $\epsilon$ -exact approximation algorithm that reduces it to  $\mathcal{O}(prN + npr^2M)$  where the constants p and n depends on the precision  $\epsilon$  and the bandwidth h.
- For example for N = M = 409,600 points while the direct evaluation of the density derivative takes around 12.76 hours the fast evaluation requires only 65 seconds with an error of around  $\epsilon = 10^{-12}$ .

#### Notion of $\epsilon - exact$ approximation

For any given  $\epsilon > 0$  the algorithm computes an approximation  $\hat{G}_r(y_j)$  such that

$$\left|\frac{\widehat{G}_r(y_j) - G_r(y_j)}{Q}\right| \le \epsilon,$$

where  $Q = \sum_{i=1}^{N} |q_i|$ .

We call  $\hat{G}_r(y_j)$  an  $\epsilon - exact$  approximation to  $G_r(y_j)$ .

- $\epsilon$  can be arbitrarily small.
- For machine precision accuracy there is no difference between the direct and the fast methods.

# Algorithm

- The fast algorithm is based on separating the  $x_i$  and  $y_j$ .
- The Gaussian is factorized via Taylor's series.
- Only first few terms are retained.
- Need to derive good error bounds to decide how many terms to retain to achieve a desired error.
- The Hermite is factorized via the binomial theorem.

# Factorization of the Gaussian

$$e^{-\|y_j - x_i\|^2 / h_2^2} = \sum_{k=0}^{p-1} \frac{2^k}{k!} \left[ e^{-\|x_i - x_*\|^2 / h_2^2} \left( \frac{x_i - x_*}{h_2} \right)^k \right] \left[ e^{-\|y_j - x_*\|^2 / h_2^2} \left( \frac{y_j - x_*}{h_2} \right)^k \right] + error_p.$$

where,

$$error_p \leq \frac{2^p}{p!} \left(\frac{\|x_i - x_*\|}{h_2}\right)^p \left(\frac{\|y_j - x_*\|}{h_2}\right)^p e^{-(\|x_i - x_*\| - \|y_j - x_*\|)^2/h_2^2}.$$

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# **Factorization of the Hermite**

$$H_r\left(\frac{y_j - x_i}{h_1}\right) = \sum_{l=0}^{\lfloor r/2 \rfloor} \sum_{m=0}^{r-2l} a_{lm} \left(\frac{x_i - x_*}{h_1}\right)^m \left(\frac{y_j - x_*}{h_1}\right)^{r-2l-m}$$

where,

$$a_{lm} = \frac{(-1)^{l+m} r!}{2^{l} l! m! (r-2l-m)!}.$$

#### Ignore the error terms and regroup

$$\widehat{G}_{r}(y_{j}) = \sum_{k=0}^{p-1} \sum_{l=0}^{\lfloor r/2 \rfloor} \sum_{m=0}^{r-2l} a_{lm} B_{km} e^{-\|y_{j}-x_{*}\|^{2}/h_{2}^{2}} \left(\frac{y_{j}-x_{*}}{h_{2}}\right)^{k} \left(\frac{y_{j}-x_{*}}{h_{1}}\right)^{r-2l-m}$$

where

$$B_{km} = \frac{2^k}{k!} \sum_{i=1}^N q_i e^{-\|x_i - x_*\|^2 / h_2^2} \left(\frac{x_i - x_*}{h_2}\right)^k \left(\frac{x_i - x_*}{h_1}\right)^m$$

- The coefficients  $B_{km}$  can be evaluated separately in  $\mathcal{O}(prN)$ .
- Evaluation of  $\hat{G}_r(y_j)$  at M points is  $\mathcal{O}(pr^2M)$ .
- Hence the computational complexity has reduced from the quadratic  $\mathcal{O}(rNM)$  to the linear  $\mathcal{O}(prN + pr^2M)$ .

## Other tricks

- Space subdivision.
- Rapid decay of the Gaussian.
- Choosing p based on tight error bounds.

# **Numerical Experiments**

- Algorithm programmed in C++ with MATLAB bindings.
- Experiments run on 2.4 GHz processor with 2 GB RAM.
- Source and target points uniformly distributed in the unit interval.

As a function of N [M = N h = 0.1 r = 4]



Linear in N.

**Precision Vs Speedup**  $[M = N = 50,000 \ h = 0.1 \ r = 4]$ 



Better speedup for reduced precision.

As a function of bandwidth h [M = N = 50,000 r = 4]



Better speedups for large bandwidths.

As a function of r [ $M = N = 50,000 \ h = 0.1$ ]



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## Speedup for bandwidth estimation

- Used the solve-the-equation plug-in method of Jones at.al (1996) \*.
- We demonstrate the speedup achieved on the mixture of normal densities used by Marron and Wand (1992).
  - A typical representative of the densities likely to be encountered in real data situations.
- The absolute relative error is defined as  $\frac{|h_{direct} h_{fast}|}{h_{direct}}$ .
- For 50,000 points we obtained speedups in the range 65 to 105 with the absolute relative error of the order  $10^{-5}$  to  $10^{-7}$ .

\*Sheather, S. and Jones, M. 1991. A reliable data-based bandwidth selection method for kernel density estimation. Journal of Royal Statistical Society Series B 53, 683-690.

#### Marron Wand normal mixtures \*



\*Marron, J. S. and Wand, M. P. 1992. Exact mean integrated squared error. The Annals of Statistics 20, 2, 712-736.

# Speedup for Marron Wand normal mixtures

	$h_{direct}$	$h_{fast}$	$T_{direct}$ (sec)	$T_{fast}$ (sec)	Speedup	Rel. Err.
1	0.122213	0.122215	4182.29	64.28	65.06	1.37e-005
2	0.082591	0.082592	5061.42	77.30	65.48	1.38e-005
3	0.020543	0.020543	8523.26	101.62	83.87	1.53e-006
4	0.020621	0.020621	7825.72	105.88	73.91	1.81e-006
5	0.012881	0.012881	6543.52	91.11	71.82	5.34e-006
6	0.098301	0.098303	5023.06	76.18	65.93	1.62e-005
7	0.092240	0.092240	5918.19	88.61	66.79	6.34e-006
8	0.074698	0.074699	5912.97	90.74	65.16	1.40e-005
9	0.081301	0.081302	6440.66	89.91	71.63	1.17e-005
10	0.024326	0.024326	7186.07	106.17	67.69	1.84e-006
11	0.086831	0.086832	5912.23	90.45	65.36	1.71e-005
12	0.032492	0.032493	8310.90	119.02	69.83	3.83e-006
13	0.045797	0.045797	6824.59	104.79	65.13	4.41e-006
14	0.027573	0.027573	10485.48	111.54	94.01	1.18e-006
15	0.023096	0.023096	11797.34	112.57	104.80	7.05e-007

# Projection pursuit

The idea of projection pursuit is to search for projections from highto low-dimensional space that are most *interesting* \*.

- 1. Given N data points in a d dimensional space project each data point onto the direction vector  $a \in \mathbf{R}^d$ , i.e.,  $z_i = a^T x_i$ .
- 2. Compute the univariate nonparametric kernel density estimate,  $\hat{p}$ , of the projected points  $z_i$ .
- 3. Compute the projection index I(a) based on the density estimate.
- 4. Locally optimize over the the choice of a, to get the *most inter*esting projection of the data.

\*Huber, P. J. 1985. Projection pursuit. The Annals of Statistics 13, 435-475.

#### **Projection index**

The projection index is designed to reveal specific structure in the data, like clusters, outliers, or smooth manifolds.

The entropy index based on Rényi's order-1 entropy is given by

$$I(a) = \int p(z) \log p(z) dz.$$

The density of zero mean and unit variance which uniquely minimizes this is the standard normal density.

Thus the projection index finds the direction which is most nonnormal.

#### Speedup

The computational burden is reduced in the following three instances.

- 1. Computation of the kernel density estimate (i.e. use the fast method with r = 0).
- 2. Estimation of the optimal bandwidth.
- 3. Computation of the first derivative of the kernel density estimate, which is required in the optimization procedure.

#### Projection pursuit on a image

(a)









The entire procedure took 15 minutes while that using the direct method takes around 7.5 hours.

#### Conclusions

- Fast  $\epsilon exact$  algorithm for kernel density derivative estimation which reduced the computational complexity from  $O(N^2)$ to O(N).
- We demonstrated the speedup achieved for optimal bandwidth estimation.
- We demonstrated how to potentially speedup the projection pursuit algorithm.

#### Software

- The code is available for academic use.
- www.cs.umd.edu/~vikas
- A detailed version of this paper is available as a TR \*.

\*Very fast optimal bandwidth selection for univariate kernel density estimation. Vikas C. Raykar and R. Duraiswami, CS-TR-4774, Department of computer science, University of Maryland, Collegepark.

#### **Related work**

- FFT \*, FGT <sup>†</sup>, IFGT <sup>‡</sup>, dual-tree <sup>§</sup>. All the above methods are designed to specifically accelerate the KDE.
- The main contribution of this paper is to accelerate the kernel density derivative estimate with an emphasis to solve the optimal bandwidth problem. The case of KDE arises as a special case of r = 0, i.e., the zero order density derivative.

\*Silverman, B. W. 1982. Algorithm AS 176: Kernel density estimation using the fast Fourier transform. Journal of Royal Statistical society Series C: Applied statistics 31, 1, 93-99.

<sup>†</sup>Greengard, L. and Strain, J. 1991. The fast Gauss transform. SIAM Journal of Scientic and Statistical Computing 12, 1, 79-94.

<sup>‡</sup>Yang, C., Duraiswami, R., Gumerov, N., and Davis, L. 2003. Improved fast Gauss transform and efficient kernel density estimation. In IEEE International Conference on Computer Vision. 464-471.

<sup>§</sup>Gray, A. G. and Moore, A. W. 2003. Nonparametric density estimation: Toward computational tractability. In SIAM International conference on Data Mining.